

XLV Konferencja Statystyka Matematyczna

Adaptive Bayesian SLOPE — High-dimensional Model Selection with Missing Values

Wei Jiang
CMAP, Ecole Polytechnique

Malgorzata Bogdan, Julie Josse, Blazej Miasojedow, Veronika Rockova

3rd Dec.



Motivation: Paris Hospital

- *Traumabase*® data:
20000 major trauma patients \times 250 measurements.

Accident type	Age	Sex	Blood pressure	Lactate	Temperature	Platelet (G/L)
Falling	50	M	140	NA	35.6	150
Fire	28	F	NA	4.8	36.7	250
Knife	30	M	120	1.2	NA	270
Traffic accident	23	M	110	3.6	35.8	170
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Traffic accident	58	F	150	NA	38.2	400

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Develop models to help emergency doctors make decisions.
Measurements $\xrightarrow{\text{Predict}}$ Platelet $\Rightarrow X \xrightarrow{\text{Regression}} y$
- **Challenge :**
How to **select** relevant measurements with **missing values**?



Model selection in high-dimension

Linear regression model: $y = X\beta + \varepsilon$,

- $y = (y_i)$: vector of response of length n
- $X = (X_{ij})$: a standardized design matrix of dimension $n \times p$
- $\beta = (\beta_j)$: regression coefficient of length p
- $\varepsilon \sim \mathcal{N}(0, \sigma^2 I_n)$

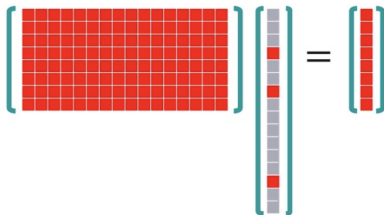
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Assumptions:

- high-dimension: p large (including $p \geq n$)
- β is **sparse** with $k < n$ nonzero coefficients


$$\begin{bmatrix} \text{Red Grid} \end{bmatrix} \begin{bmatrix} \text{Tall Vector with sparse red cells} \end{bmatrix} = \begin{bmatrix} \text{Tall Vector with all red cells} \end{bmatrix}$$

l_1 penalization methods

- LASSO (Tibshirani, 1996)

$$\hat{\beta}_{LASSO} = \arg \min_{\beta \in \mathbb{R}^p} \frac{1}{2} \|y - X\beta\|^2 + \lambda \|\beta\|_1,$$

detects important variables with high probability but includes many **false positives**.

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where $\lambda_1 \geq \lambda_2 \geq \dots \geq \lambda_p \geq 0$ and $|\beta|_{(1)} \geq |\beta|_{(2)} \geq \dots \geq |\beta|_{(p)}$.

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To control **False Discovery Rate (FDR)** at level q :

$\lambda_{BH}(j) = \phi^{-1}(1 - q_j)$, $q_j = \frac{jq}{2p}$, $X^T X = I$, then

$$FDR = \mathbb{E} \left[\frac{\# \text{False rejections}}{\# \text{Rejections}} \right] \leq q$$



Problem: λ for SLOPE leading to FDR control are typically large.
SLOPE often returns **an inconsistent estimation.**

\Rightarrow improve?



Bayesian SLOPE

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SLOPE estimate = MAP of a Bayesian regression with SLOPE prior.

$$\hat{\beta}_{SLOPE} = \arg \max_{\beta} p(y \mid X, \beta, \sigma^2; \lambda) \propto p(y \mid X, \beta) p(\beta \mid \sigma^2; \lambda)$$

where the SLOPE prior:

$$p(\beta \mid \sigma^2; \lambda) \propto \prod_{j=1}^p \exp \left(-\frac{1}{\sigma} \lambda_j |\beta|_{(j)} \right)$$

Adaptive Bayesian SLOPE

We propose an adaptive version of Bayesian SLOPE (ABSLOPE), with the prior for β as

$$p(\beta \mid \gamma, c, \sigma^2; \lambda) \propto c^{\sum_{j=1}^p \mathbb{I}(\gamma_j=1)} \prod_j \exp \left\{ -w_j |\beta_j| \frac{1}{\sigma} \lambda_{r(w_{\beta,j})} \right\},$$

Interpretation of the model:

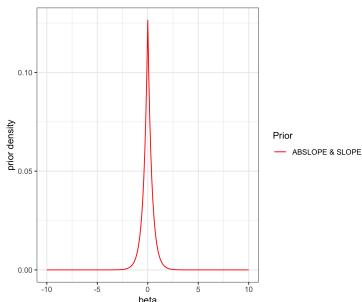
- β_j is large enough \Rightarrow **true signal**; 0 \Rightarrow noise.
- $\gamma_j \in \{0, 1\}$ signal indicator. $\gamma_j | \theta \sim \text{Bernoulli}(\theta)$ and θ the **sparsity**.
- $c \in [0, 1]$: the inverse of **average signal magnitude**.
- $W = \text{diag}(w_1, w_2, \dots, w_p)$ and its diagonal element:

$$w_j = c\gamma_j + (1 - \gamma_j) = \begin{cases} c, & \gamma_j = 1 \\ 1, & \gamma_j = 0 \end{cases}.$$

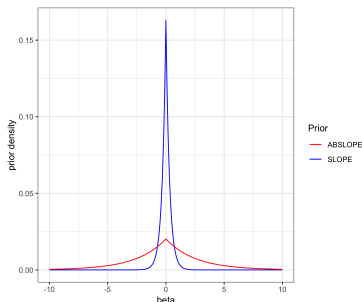
Adaptive Bayesian SLOPE

Advantage of introducing W :

- when $\gamma_j = 0, w_j = 1$, i.e., the null variables are treated with the regular SLOPE penalty
- when $\gamma_j = 1, w_j = c < 1$, i.e., **smaller penalty** $\lambda_{r(W\beta, j)}$ for true predictors than the regular SLOPE one



(a) Null β



(b) Non-null β

Figure: comparison of SLOPE prior and ABSLOPE prior



Model selection with missing values

Decomposition: $X = (X_{\text{obs}}, X_{\text{mis}})$

Pattern: matrix M with $M_{ij} = \begin{cases} 1, & \text{if } X_{ij} \text{ is observed} \\ 0, & \text{otherwise} \end{cases}$

Assumption 1: Missing at random (MAR)

$p(M \mid X_{\text{obs}}, X_{\text{mis}}) = p(M \mid X_{\text{obs}}) \Rightarrow$ ignorable missing patterns
e.g. People at **older age** didn't tell his **income** at larger probability.

Assumption 2: Distribution of covariates

$X_i \sim_{\text{i.i.d.}} \mathcal{N}_p(\mu, \Sigma), \quad i = 1, \dots, n.$

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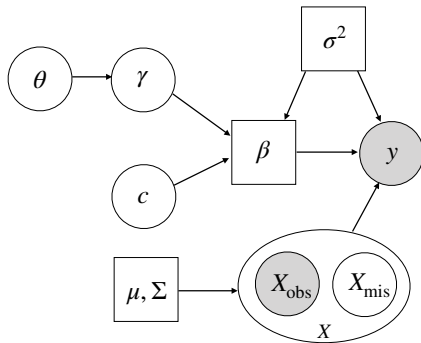
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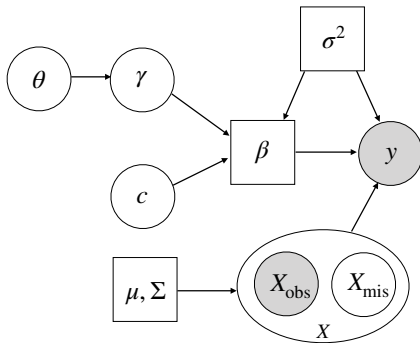
Problem: With NA, only a few methods are available to select a model, and their performances are limited. For example,

- (Claeskens and Consentino, 2008) adapts AIC to missing values \Rightarrow Impossible to deal with high dimensional analysis.
- (Loh and Wainwright, 2012) LASSO with NA
 \Rightarrow Non-convex optimization; requires to know bound of $\|\beta\|_1$
 \Rightarrow difficult in practice

ABSLOPE with missingness: Summary



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$$\begin{aligned}\ell_{\text{comp}} &= \log \mathbf{p}(y, X, \gamma, c; \beta, \theta, \sigma^2) + \text{pen}(\beta) \\ &= \log \{ \mathbf{p}(X; \mu, \Sigma) \mathbf{p}(y \mid X; \beta, \sigma^2) \mathbf{p}(\gamma; \theta) \mathbf{p}(c) \} + \text{pen}(\beta)\end{aligned}$$

Objective: Maximize $\ell_{\text{obs}} = \iiint \ell_{\text{comp}} dX_{\text{mis}} dc d\theta d\gamma$.



EM algorithm

- *E step*: evaluate

$$Q^t = \mathbb{E}(\ell_{\text{comp}}) \quad \text{wrt} \quad p(X_{\text{mis}}, \gamma, c, \theta \mid y, X_{\text{obs}}, \beta^t, \sigma^t, \mu^t, \Sigma^t).$$

- *M step*: update

$$\beta^t, \sigma^t, \mu^t, \Sigma^t = \arg \max Q^t$$

Problem: The function Q is not tractable. \Rightarrow

- ① Monte Carlo EM ? (Wei and Tanner 1990)



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- ① ~~Monte Carlo EM?~~

Expensive to generate a large number of samples.

- ② Stochastic Approximation EM (book, Lavielle 2014)
 - One sample in each iteration;

Adapted SAEM algorithm

- *E step:*

$$Q^t = \mathbb{E}(\ell_{\text{comp}}) \quad \text{wrt} \quad p(X_{\text{mis}}, \gamma, c, \theta \mid y, X_{\text{obs}}, \beta^t, \sigma^t, \mu^t, \Sigma^t).$$

- *Simulation:* draw one sample $(X_{\text{mis}}^t, \gamma^t, c^t, \theta^t)$ from

$$p(X_{\text{mis}}, \gamma, c, \theta \mid y, X_{\text{obs}}, \beta^{t-1}, \sigma^{t-1}, \mu^{t-1}, \Sigma^{t-1});$$

[**Gibbs sampling**]

- *Stochastic approximation:* update function Q with

$$Q^t = Q^{t-1} + \eta_t \left(\ell_{\text{comp}} \Big|_{X_{\text{mis}}^t, \gamma^t, c^t, \theta^t} - Q^{t-1} \right).$$

- *M step:* $\beta^{t+1}, \sigma^{t+1}, \mu^{t+1}, \Sigma^{t+1} = \arg \max Q^{t+1}.$

[**Proximal gradient descent, Shrinkage of covariance**]

Details of initialization, generating samples and optimization are in the draft
([arXiv:1909.06631](https://arxiv.org/abs/1909.06631))



R package: ABSLOPE

Install package:

```
library(devtools)  
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Main algorithm:

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lambda = create_lambda_bhq(ncol(X),fdr=0.10)  
list.res = ABSLOPE(X, y, lambda)
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A fast and simplified algorithm (Rcpp):

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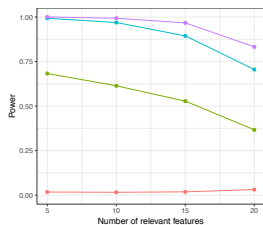
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Values:

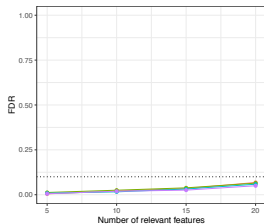
```
list.res$beta  
list.res$gamma
```


Simulation study (200 rep. \Rightarrow average)

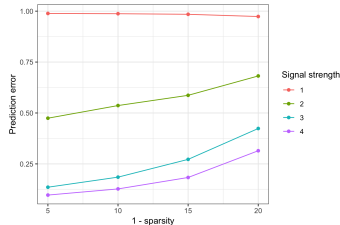
$n = p = 100$, no correlation and 10% missingness



(a) Power



(b) FDR

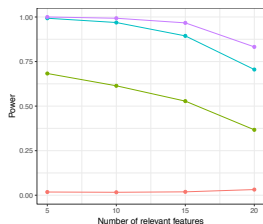


(c) Prediction error

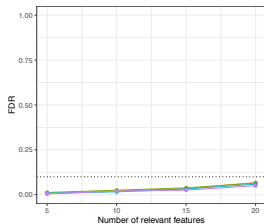


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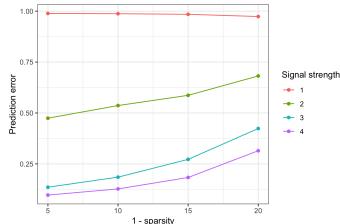
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(g) Power

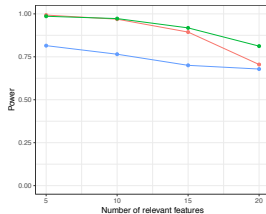


(h) FDR

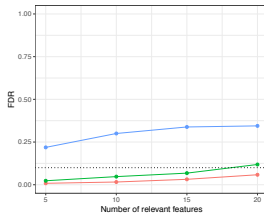


(i) Prediction error

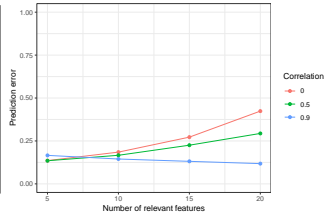
$n = p = 100$, with 10% missingness and strong signal



(j) Power



(k) FDR



(l) Prediction error



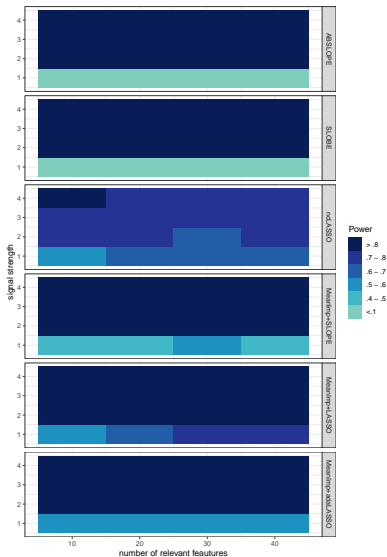
Method comparison

- **ABSLOPE** and **SLOBE**
- **ncLASSO**: non convex LASSO (Loh and Wainwright, 2012)
- **MeanImp + SLOPE**: Mean imputation followed by SLOPE with known σ
- **MeanImp + LASSO**: Mean imputation followed by LASSO, with λ tuned by cross validation
- **MeanImp + adaLASSO**: Mean imputation followed by adaptive LASSO (Zou, 2006)

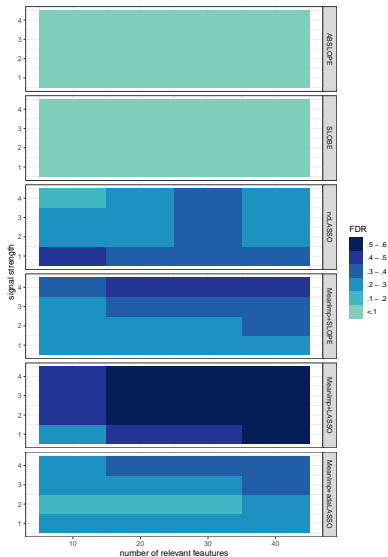
In the SLOPE type methods, $\lambda = \text{BH sequence}$ which controls the FDR at level **0.1**

Method comparison (200 rep. \Rightarrow average)

500 \times 500 dataset, 10% missingness, with correlation



(m) Power

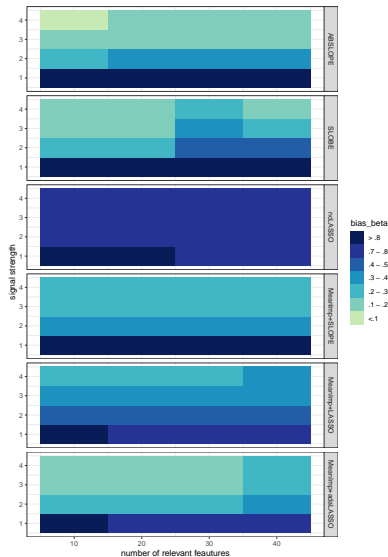


(n) FDR

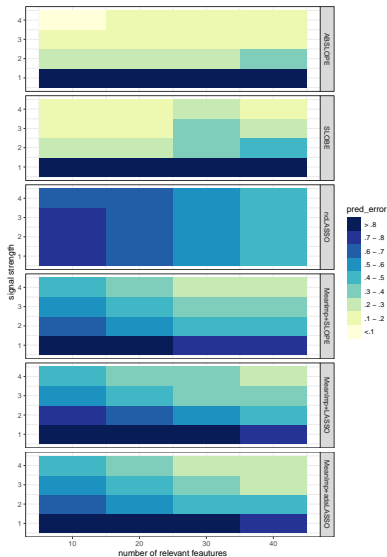


Method comparison (200 rep. \Rightarrow average)

500 \times 500 dataset, 10% missingness, with correlation



(a) Bias of β



(b) Prediction error



Computational cost

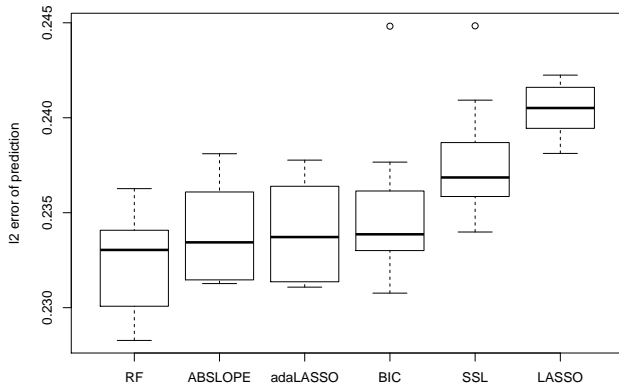
Execution time (seconds) for one simulation	$n = p = 100$			$n = p = 500$		
	min	mean	max	min	mean	max
ABSLOPE	12.83	14.33	20.98	646.53	696.09	975.73
SLOBE	0.31	0.34	0.66	14.23	15.07	29.52
ncLASSO	16.38	20.89	51.35	91.90	100.71	171.00
MeanImp + SLOPE	0.01	0.02	0.09	0.24	0.28	0.53
MeanImp + LASSO	0.10	0.14	0.32	1.75	1.85	3.06

[**Fast implementation:** Parallel computing + Rcpp (C++)]

More on the real data...

TraumaBase: Measurements $\xrightarrow{\text{Predict}}$ Platelet

Cross-validation: random splits to training and test sets $\times 10$



- Comparable to random forest
- Interpretable model selection and estimation results



Conclusion & Future research

Conclusion:

- ABSLOPE penalizes larger coefficients more stringently to **control FDR**, meanwhile it applies a weighting matrix to **improve the estimation**;
- Modeling in a Bayesian framework gives detailed structure of predictors as **sparsity** and **signal strength**;
- Simulation study shows that ABSLOPE is competitive to other methods in terms of power, FDR and prediction error.

Future research:

- Consider categorical or mixed data
- Deal with other missing mechanisms
- Application on genetic dataset

Thank you!
Dziękuję
Merci

